Dynamical dark energy in light of the latest observations

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A flat Friedmann-Robertson-Walker universe dominated by a cosmological constant (Λ) and cold dark matter (CDM) has been the working model preferred by cosmologists since the discovery of cosmic acceleration^{1,2}. However, tensions of various degrees of significance are known to be present among existing datasets within the Λ CDM framework³⁻¹¹. In particular, the Lyman- α forest measurement of the baryon acoustic oscillations (BAO) by the Baryon Oscillation Spectroscopic Survey³ prefers a smaller value of the matter density fraction $\Omega_{\rm M}$ than that preferred by cosmic microwave background (CMB). Also, the recently measured value of the Hubble constant, $H_0 = 73.24 \pm 1.74 \,\mathrm{km}\,\mathrm{s}^{-1}\,\mathrm{Mpc}^{-1}$ (ref. ¹²), is 3.4σ higher than the $66.93 \pm 0.62 \,\mathrm{km}\,\mathrm{s}^{-1}\mathrm{Mpc}^{-1}$ inferred from the Planck CMB data7. In this work, we investigate whether these tensions can be interpreted as evidence for a non-constant dynamical dark energy. Using the Kullback-Leibler divergence¹³ to quantify the tension between datasets, we find that the tensions are relieved by an evolving dark energy, with the dynamical dark energy model preferred at a 3.5 σ significance level based on the improvement in the fit alone. While, at present, the Bayesian evidence for the dynamical dark energy is insufficient to favour it over Λ CDM, we show that, if the current best-fit dark energy happened to be the true model, it would be decisively detected by the upcoming Dark Energy Spectroscopic Instrument survey¹⁴.

The observational datasets considered in this work include the latest cosmic microwave background (CMB) temperature and polarization anisotropy spectra, the supernovae (SNe) luminosity distance data, the baryon acoustic oscillations (BAO) angular diameter distance data from the clustering of galaxies (gBAO) and from the Lyman- α forest (Ly α FB), the estimate of the current value of the Hubble constant (H_0), the measurements of the Hubble parameter H(z) at redshifts z using the relative age of old and passively evolving galaxies (OHD), the three-dimensional galaxy power spectra and the two-dimensional weak-lensing shear angular power spectra. Further details about the datasets and associated systematic effects can be found in the Methods.

The Kullback–Leibler (KL) divergence, also known as relative entropy, quantifies the proximity of two probability density functions (PDFs). Rather than focusing on particular model parameters, it is designed to compare the overall concordance of datasets within a given model. We use the difference between the actual and the expected KL divergence, called 'Surprise'¹⁵, as a measure of tension between datasets. Rather than comparing the PDFs for the Λ cold dark matter model (Λ CDM, where Λ is the cosmological constant) parameters for every pair of datasets, we take the combined dataset, ALL16 (see Supplementary Table 1), and find the derived PDFs for the angular diameter distance $D_{\Lambda}(z)$ and H(z) at redshifts corresponding to the available data. We then compute the KL divergence between the derived PDFs and the directly observed $D_{\Lambda}(z)$

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and H(z) from H_0 , SNe, OHD, gBAO and Ly α FB, and evaluate the corresponding Surprise and the standard deviation (see Methods for details). Results are shown as cyan bars in Fig. 1a. They indicate that the H_0 , Ly α FB and SNe measurements are in tension with the combined dataset. Introducing tension T as the number of standard deviations by which Surprise is greater than zero, we find values of T = 4.4, 3.5 and 1.7 for the H_0 , Ly α FB and SNe measurements, respectively (shown in Fig. 1b), with the first two values signalling significant tension.

Next, we check whether the tension within the Λ CDM model can be interpreted as evidence for a dynamical dark energy. The dynamics of dark energy can be probed in terms of its equation of state w, which is equal to -1 for Λ , but is different in dynamical dark energy models, where it will generally be a function of redshift z. Commonly considered alternatives to Λ are a model with a constant w (wCDM) and one in which w is a linear function of the scale factor $(w_0 w_a CDM)^{16}$. We allow for a general evolution of the dark energy equation of state and use the correlated prior probability distribution method¹⁷ (hereafter, the prior probability distribution is simply called the 'prior') to perform a Bayesian nonparametric reconstruction of w(z) using the Markov chain Monte Carlo method with other cosmological parameters marginalized over (see Methods for details). Figure 2 presents the reconstructed w(z), along with the 68% confidence-level uncertainty, shown with a light blue band, derived from the combined dataset ALL16. Table 1 shows the change in χ^2 relative to Λ CDM for each individual dataset for the best-fit w(z)CDM model derived from ALL16. Overall, the χ^2 is improved by -12.3, which can be interpreted as the reconstructed dynamical dark energy model being preferred at 3.5 σ . The reconstructed dark energy equation of state evolves with time and crosses the -1 boundary, which is prohibited in singlefield minimally coupled quintessence models¹⁸, but can be realized in models with multiple scalar fields, such as Quintom¹⁹, or if the dark energy field mediates a new force between matter particles²⁰. In the latter case, which is commonly classified as modified gravity, it is generic for the effective dark energy equation of state to be close to -1 around z=0, but to evolve towards more negative values at intermediate redshifts, before eventually approaching 0 during matter domination. Such dynamics would be consistent with our reconstruction and could be tested in the future when BAO measurements at higher redshifts become available. In addition to the reconstruction from the combined ALL16 dataset presented in

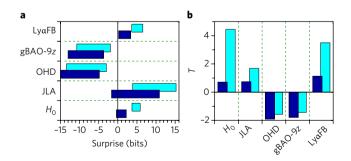


Fig. 1 The tension among different datasets in Λ CDM and w(z)CDM universes. **a**, The Surprise between the PDFs for $D_A(z)$ and H(z) derived from the best-fit model using the combined dataset of ALL16, and the directly observed $D_A(z)$ and H(z) from H_O , JLA (the JLA sample of SNe), OHD, gBAO-9z (gBAO measurements at nine effective redshifts) and LyαFB, respectively (see Methods for detailed explanation and references for data used). The cyan horizontal bars indicate the 68% confidence-level range of Surprise in Λ CDM, while the dark blue bars correspond to that of w(z)CDM. **b**, The corresponding values of tension T, defined as Surprise divided by its standard deviation, shown using the same colour scheme as in **a**.

Fig. 2, we present reconstructions derived from ten different data combinations in Supplementary Fig. 1.

The results for tension between datasets, re-evaluated for the ALL16 best-fit w(z)CDM model, are shown as dark blue bars in Fig. 1. We find T=0.7, 1.1 and 0.7 for H_0 , Ly α FB and SNe, respectively, indicating that tensions that existed in the Λ CDM model are significantly released within w(z)CDM. A plot of the relevant data points along with the best-fit predictions from the Λ CDM and the w(z)CDM models are provided in Supplementary Fig. 2.

With a large number of additional w-bin parameters, one may be concerned that the improvement in the fit is achieved by w(z)CDM at the cost of a huge increase in the parameter space. However, correlations between the w-bins induced by the prior constrain most of that freedom. One way to estimate the effective number of additional degrees of freedom is to perform a principal-component analysis of the posterior covariance matrix of the w-bin parameters and compare it with that of the prior. Using this method, explained in detail in the Methods section, we find that our w(z)CDM model effectively has only four additional degrees of freedom compared with Λ CDM. We note that the demonstration that ALL16 is capable of constraining four principal components of w(z) is one of the interesting results of this work.

It is interesting to compare w(z) reconstructed from ALL16 with that obtained by Zhao et al.²¹ using the same prior but a different dataset, which we call ALL12 (a comparison of the ALL16 and the ALL12 datasets is provided in Supplementary Table 2). ALL16 contains about 40% more SNe than ALL12, primarily provided by the Sloan Digital Sky Survey (SDSS)-II. Moreover, in ALL12, the BAO measurement derived from the Baryon Oscillation Spectroscopic Survey (BOSS) DR9 sample²² was at a single effective redshift, whereas in ALL16 it is tomographic at nine redshifts from BOSS DR12 (ref. ²³), which contains four times more galaxies than DR9. In addition, ALL16 includes a high-redshift Ly α FB measurement,

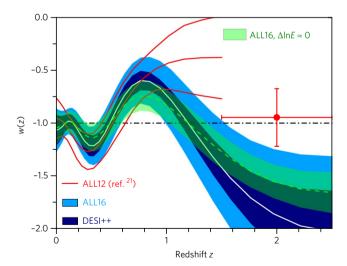


Fig. 2 | Reconstructed evolution history of the dark energy equation of state compared with the 2012 result and the forecasted uncertainty from future data. The mean (white solid) and the 68% confidence level (CL) uncertainty (light blue band) of the w(z) reconstructed from ALL16 compared with the ALL12 w(z) reconstructed by Zhao et al. (ref. ²¹) (red lines show the mean and the 68% CL band). The red point with 68% CL error bars is the value of w(z) at z=2 'predicted' by the ALL12 reconstruction. The dark blue band around the ALL16 reconstruction is the forecasted 68% CL uncertainty from DESI++. The green dashed curve and the light green band show the mean and the 68% CL of w(z), respectively, reconstructed from ALL16 using a different prior strength ($\sigma_D=0.4$), for which the Bayesian evidence is equal to that of ΛCDM. See the text for details.

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Table 1 The improved χ^2 with respect to that of the Λ CDM model									
	P15	JLA	gBAO	P(k)	WL	H _o	LyαFB	OHD	
$\chi^2_{w(z)CDM} - \chi^2_{\Lambda CDM}$	-0.7	-1.6	-2.8	+1.1	-0.1	-2.9	-3.7	-2.3	
$\chi_{\text{wCDM}}^2 - \chi_{\Lambda\text{CDM}}^2$	0.0	+0.5	+0.7	+0.4	+0.2	-2.9	-0.2	0.0	
$\chi^2_{w_0w_aCDM} - \chi^2_{\LambdaCDM}$	-0.7	+0.4	+0.9	+0.5	+0.4	-2.7	-0.3	0.0	

The changes in χ² of individual datasets between the ALL16 best-fit w(z)CDM, wCDM and w₀w₀CDM models and the ΛCDM model.

which was not available in 2012. This helps to constrain w(z) at redshifts where the SNe constraints are weak. The new H_0 measurement in 2016 (ref. ¹²) is consistent with that in 2009 (ref. ²⁴), with the error bar halved. Comparing measurements of the expansion rate and the cosmic distances in ALL12 and ALL16, we find that those in ALL16 offer information at a greater number of redshift values, and with a greater signal-to-noise ratio (see Supplementary Fig. 4 for a visual comparison). Quantitatively, the signal-to-noise ratio in measurements of H(z), $D_{\rm A}(z)$ and $d_{\rm L}(z)$ in ALL16 is larger by 80%, 260% and 90%, respectively, than in those of the ALL12 dataset, where $d_{\rm L}(z)$ is the luminosity distance. The Planck 2015 CMB data is also much more informative than the Wilkinson Microwave Anisotropy Probe seven-year release²⁵, thanks to a higher angular resolution of the temperature and polarization maps, and lower levels of statistical uncertainties.

Overall, ALL16 is more constraining due to a significant level of new and independent information in ALL16 compared with ALL12: the effective number of w(z) degrees of freedom constrained by ALL12 was three, compared with four constrained by ALL16. Figure 2 compares the two results and shows that they are highly consistent. We quantify the agreement by evaluating the dot product of the \hat{w} vectors from the two reconstructions (the vectors are normalized so that a dot product is unity if they are identical) and find that $\hat{w}_{\text{ALL12}} \cdot \hat{w}_{\text{ALL16}} = 0.94 \pm 0.02$. We also evaluate the tension T between the two reconstruction results and find that T = -1.1. This indicates an excellent alignment of the two results. This agreement, and the raised significance of an evolving w(z) from 2.5σ to 3.5σ confidence levels with more advanced observations, suggest the possibility of revealing the dynamics of dark energy at a much more statistically significant level in the near future, as we will present later.

To check whether the improvement in the fit warrants introducing additional effective degrees of freedom, we evaluate and compare the Bayesian evidence, $E \equiv \int d\theta \mathcal{L}(D|\theta) P(\theta)$, for the Λ CDM and the w(z)CDM models. The Bayes factors, which are the differences in ln between the two models, are shown in Table 2. The Bayes factors for both the ALL12 and the ALL16 dark energy models are negative, indicating that Λ CDM is favoured by this criterion. However, our forecast for a future dataset, DESI++, comprised of BAO measurements from the Dark Energy Spectroscopic Instrument (DESI)¹⁴, around 4,000 SNe luminosity distances from future surveys²⁶ and the CMB (assuming Planck sensitivity), predicts a Bayes factor of 11.3 \pm 0.3 if the ALL16 w(z)CDM reconstruction happened to be the true model, which would be decisive according to the Jeffreys scale.

 Table 2 | The significance of $w \neq -1$ and Bayes factors

 ALL12
 ALL16
 DESI++

	ALLIZ	ALLIO	DESI++
Signal-to-noise ratio	2.5σ	3.5σ	6.4σ
$ln(E_{w(z)CDM}/E_{\Lambda CDM})$	-6.7 ± 0.3	-3.3 ± 0.3	11.3 ± 0.3

The statistical significance in terms of the standard deviation σ , that is, the signal-to-noise ratio defined as $\sqrt{\Delta\chi^2}$, of w(z) deviating from -1 based on the improvement in the fit alone, and the Bayes factor $\Delta \ln E$ between the w(z)CDM models and the Λ CDM model, derived from the ALL12 and the ALL16 datasets respectively, along with the forecast for DESI++.

One may ask how much the evidence for dark energy depends on the particular choice of the prior parameters. In principle, the choice of the smoothing scale should be guided by theory. The value used by Zhao et al.²¹ and in this work, $a_c = 0.06$, where a_c describes the typical smoothing scale, is a timescale conservatively chosen to be sufficiently small so as not to bias reconstructions of w(z) expected in quintessence dark energy models¹⁷. For the inference to be conclusive, the evidence for a dynamical dark energy should be strong over a wide range of the prior parameters. Therefore, we vary the strength of our prior by adjusting σ_D , which is a parameter added to the diagonal of the inverse of the prior covariance matrix, and examine how the significance of the dynamical dark energy detection, as well as the Bayesian evidence, change with the variation of σ_D . As shown in Fig. 3, and with additional details given in the Methods section, we find that neither ALL12 nor ALL16 provides evidence for a dynamical dark energy over the considered wide range of prior strengths. However, the Bayes factor for ALL16 is generally much

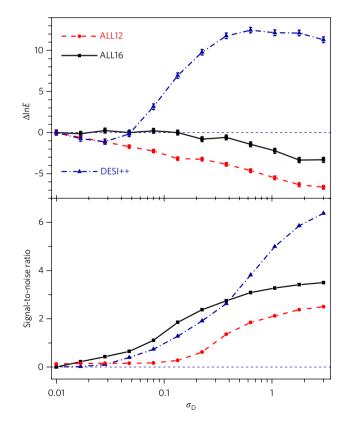


Fig. 3 | The Bayes factor and the significance level of $w \neq -1$ for various correlated priors for current and future data. The Bayes factor with 68% confidence-level error bars (upper panel) and the statistical significance (lower panel) of dynamical dark energy derived from the 2012 data (ALL12; red dashed)²¹, current data (ALL16; black solid) and future data (DESI++; blue dot-dashed)¹⁴. Error bars show standard deviation.

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Table 3 Results for extended models							
	M ₁	M_2	M_3	M_4	M ₅		
Definition	$M_{\nu} = 0.06 \mathrm{eV}$	M_{ν} , $N_{\rm eff}$	A_{Lens}	wCDM	w_0w_aCDM		
Signal-to-noise ratio	3.6σ	3.4σ	3.4σ	3.5σ	3.4σ		
$\Delta \ln E$	-2.8 ± 0.3	-3.9 ± 0.3	-3.6 ± 0.3	-0.9 ± 0.3	0.7 ± 0.3		

The statistical significance, signal-to-noise ratio = $\sqrt{\Delta\chi^2}$, of the preference for w(z)CDM over Λ CDM with added parameters (M_v, M_2, M_3) , and over a model with a constant w (that is, M_s) and one with a linearly varying w(a) (that is, M_s). M_1 denotes a model with M_b =0.06 eV, M_2 contains varying M_a and N_{ell} , and M_3 includes a varying A_{Lens} . For M_b , M_2 and M_3 , the new parameters are added to both w(z) CDM and Λ CDM. The last row shows the Bayes factors. Δ In E, between w(z)CDM and the five extended models.

less negative than that for ALL12 for all prior strengths. For example, the Bayes factor inferred from ALL12 increased from -6.7 ± 0.3 to -3.3 ± 0.3 for $\sigma_{\rm D}\!=\!3$, which is the prior used in this work. In fact, for ALL16, the Bayes factor remains close to zero for $\sigma_{\rm D}\!\lesssim\!0.4$. We plot the model with $\sigma_{\rm D}\!=\!0.4$ as a light green band in Fig. 2 to demonstrate the impact of changing the prior strength, and because it is a model that has the same Bayesian evidence as $\Lambda{\rm CDM}$ while deviating from Λ at a 2.7 σ confidence level. Conversely, our forecast for DESI++ shows that, if the $w(z){\rm CDM}$ model was true, it would be decisively supported by Bayesian evidence over a wide range of prior strengths, as shown in Fig. 3.

Various ways to relieve the tension between datasets have been proposed, including allowing for additional relativistic degrees of freedom²⁷, massive neutrinos²⁷ and an interacting vacuum^{28,29}. In addition, to relieve the tension between the Λ CDM parameters required to fit the CMB temperature anisotropy spectrum at large and small scales, the Planck team introduced a parameter, $A_{\rm Lens}$, which rescales the amplitude of the weak-lensing contribution to the temperature power spectrum. In the w(z) reconstruction discussed earlier, we fixed $A_{\text{\tiny Lens}} = 1$, assumed that neutrinos are massless and set the effective number of relativistic species at the standard Λ CDM value of N_{eff} = 3.04. We have checked the effect of these parameters on the reconstructed w(z) by considering model M_1 , with neutrino mass, M_2 , fixed to 0.06 eV; model M_2 , with M_2 and $N_{\rm eff}$ added as free parameters; and model M₃, with varied $A_{\rm Lens}$. In all these cases, we find that the shape of the reconstructed w(z) and the significance of its deviation from -1 are practically the same. The inferred values of the cosmological parameters in these models are given in Supplementary Table 3, and the corresponding reconstructed w(z) values are shown in Supplementary Fig. 5. The Bayes factors for M₁, M₂ and M₃ relative to the corresponding ΛCDM models (with the same added parameters) are shown in Table 3. We also checked that neither the constant w model (M_4) nor the linear (w_0, w_a) parametrization of w (M_5) is capable of releasing the tensions between all datasets simultaneously (see Table 1). Interestingly, we find that our dark energy model with a non-parametrically reconstructed w(z) has a larger Bayes factor than that of $w_0 w_a$ CDM despite the latter having only two parameters (see results for M_4 and M_5 in Table 3).

There is always a possibility that the tensions between datasets, quantified in terms of the KL divergence in this work, are due to asyet-unknown systematic effects. However, it is intriguing that they persist with improvements in the quantity and the quality of the data 10,11,30 . If interpreted as a manifestation of dark energy dynamics, they suggest a w(z) that crosses -1 and has a shape that is representative of modified-gravity models. The commonly used (w_0, w_a) parametrization would have missed this behaviour and has a lower Bayesian evidence than the reconstructed w(z) model, despite the latter having more degrees of freedom. Thus, our results demonstrate that the current data can provide non-trivial constraints on the dark energy dynamics. It is also intriguing that the evidence for $w(z) \neq -1$, while below that of Λ CDM, has become stronger with the new independent data added since 2012, and that the ALL16 reconstruction is consistent with the ALL12 w(z). We emphasize

that we have not optimized the prior to maximize either the Bayes ratio or the statistical significance of the departure from -1, as it would be contrary to the principles of Bayesian inference. Future data have the ability to confirm conclusively the dark energy evolution inferred in this work if it happened to be the true model.

Methods

Tension calculation. The KL divergence¹³, also known as relative entropy, has been extensively used as a way of quantifying the degree of tension between different datasets within the ACDM model^{15,31–37}. Rather than focusing on particular model parameters, it is designed to compare the overall concordance of datasets within a given model. Alternative methods of quantification of the tension have been discussed in the literature^{8,38}.

The KL divergence quantifies the proximity of two PDFs, P_1 and P_2 , of a multi-dimensional random variable θ . If both P_1 and P_2 are assumed to be Gaussian¹⁵, and data are assumed to be more informative than the priors, we can write the difference between the actual and the expected KL divergence, called the 'Surprise'¹⁵, as

$$S = \frac{1}{2\ln 2} [(\theta_1 - \theta_2)^T C_1^{-1} (\theta_1 - \theta_2) - \text{tr}(C_2 C_1^{-1} + I)]$$
 (1)

where θ_1 and θ_2 are the best-fit parameter vectors, C_1 and C_2 are the covariance matrices for P_1 and P_2 , and I is the unity matrix. The standard deviation of the expected KL divergence is

$$\Sigma = \frac{1}{\sqrt{2\ln 2}} \sqrt{\operatorname{tr}(\mathcal{C}_2 \mathcal{C}_1^{-1} + I)^2}$$
 (2)

We can quantify the tension, T, between P_1 and P_2 in terms of the signal-to-noise ratio, as $T = S/\Sigma$. If $T \lesssim 1$, then P_1 and P_2 are consistent with each other 12 .

Datasets used. The datasets we consider include the Planck 2015 (P15) CMB temperature and polarization auto- and cross-angular power spectra", the Joint Light-curve Analysis SNe³⁹ (JLA), the 6dFRS (6dF)⁴⁰ and SDSS main galaxy sample⁴¹ BAO measurements, the WiggleZ galaxy power spectra⁴² in four redshift slices containing information about the BAO and redshift-space distortions (P(k)), the weak-lensing shear angular power spectra in six redshift slices from CFHTLenS⁴³ (WL), the recent estimate of H_0 obtained from local measurements of Cepheids¹² (H_0), the H(z) measurements using the relative age of old and passively evolving galaxies following a cosmic chronometer approach⁴⁴ (OHD), the BOSS DR12 'Consensus' BAO and redshift-space distortion measurements using the complete BOSS DR12 sample covering the redshift range of 0.2 < z < 0.75 at three effective redshifts (BAO-3z)⁴⁵, the BAO measurements using the same galaxy sample but at nine effective redshifts²³ (BAO-9z), and the LyαFB measurements³. A summary of datasets and data combinations used in this work is shown in Supplementary Table 1.

We account for the systematic effects in our analysis as implemented in the public likelihood codes. However, we note that there may be additional systematic effects. For example, the relative velocity between baryons and dark matter may affect the BAO distance measurements ⁴⁶. This effect is estimated to be at subpercent level for the galaxy BAO measurements of BOSS^{47,48} and is currently unknown for LyaFB. For SNe, we use the conventional χ^2 statistics for the analysis, although alternative statistics may extract more information and reduce the systematic effects for the JLA sample to some extent^{49,50}.

Non-parametric w(z) **reconstruction.** To start, w(z) is parameterized in terms of its values at discrete steps in z, or the scale factor a. Fitting a large number of uncorrelated bins would result in extremely large uncertainties and, in fact, would prevent the Markov chain Monte Carlo samples from converging because of the many degenerate directions in parameter space. Conversely, fitting only a few bins could significantly bias the result. Our approach is to introduce a prior covariance between the bins based on a specified two-point function that correlates values of w at different a, $\xi_w(|a-a'|) \equiv \langle (w[a]-w^{fid}[a])(w[a']-w^{fid}[a']) \rangle$ (where w^{fid} is the

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fiducial model), which can be taken to be of the form proposed in ref. ⁵¹, $\xi_{\text{CPZ}}(\delta a) = \xi_w(0)/(1 + [\delta a/a_c]^2), \text{ where } a \text{ is the scale factor of the Universe and } \xi_w(0) \text{ is a normalization factor set by the expected variance in the mean } w, \sigma_w^2$. As shown in ref. ¹⁷, results are largely independent of the choice of the correlation function. The prior covariance matrix C is obtained by projecting $\xi_w(|a-a'|)$ onto the discrete w-binsi^{17,51}, and the prior PDF is taken to be of Gaussian form: $\mathcal{P}_{\text{prior}}(w) \propto \exp[-(w-w^{\text{fid}})^T C^{-1}(w-w^{\text{fid}})/2]$. The reconstructed model is that which maximizes the posterior probability, which by Bayes' theorem is proportional to the likelihood of the data multiplied by the prior probability, $\mathcal{P}(w|D) \propto \mathcal{P}(D|w) \times \mathcal{P}_{\text{prior}}(w)$. Effectively, the prior results in a new contribution to the total χ^2 of a model, which penalizes models that are less smooth.

In our reconstruction of w(z), we set $a_c = 0.06$ and $\sigma_{\overline{w}} = 0.04$, which is the 'weak prior' used by Zhao et al. ²¹. To calculate the observables, we use a version of Code for Anisotropies in the Microwave Background ⁵² modified to include dark energy perturbations for an arbitrary w^{53} . We use PolyChord ⁵⁴, which is a nested sampling plug-in for CosmoMC ⁵⁵, to sample the parameter space $P \equiv (\omega_b, \omega_c, \Theta_o, \tau, n_s, A_s, w_1, \ldots, w_{30}, \mathcal{N})$ where ω_b and ω_c are the baryon and CDM densities, Θ_s is the angular size of the sound horizon at decoupling, τ is the optical depth, n_s and A_s are the spectral index and the amplitude of the primordial power spectrum, respectively, and w_1, \ldots, w_{30} denote the 30 w-bin parameters. The first 29 w-bins are uniform in $a \in (0.286, 1)$, corresponding to $z \in (0, 2.5)$, and the last wide bin covers $z \in (2.5, 1100)$. We marginalize over nuisance parameters such as the intrinsic supernova luminosity.

Principal-component analysis of w(z)**.** First, we diagonalize the posterior covariance of w-bins to find their uncorrelated linear combinations (eigenmodes), along with the eigenvalues, which quantify how well a given eigenmode is constrained56. We plot the inverse eigenvalues of the posterior covariance, ordered according to the number of nodes in the eigenmodes, in Supplementary Fig. 3a. The number of nodes is representative of the smoothness in the evolution of eigenmodes, with the first four posterior eigenmodes shown in Supplementary Fig. 3b. Next, we perform a principal-component analysis of the prior covariance matrix and plot its inverse eigenvalues alongside those of the posterior. We see that the fifth and higher number eigenvalues of the two matrices coincide, which means that they are fully determined by the prior. However, the first four inverse eigenvalues of the posterior are significantly larger than that of the prior, indicating that they are constrained primarily by the data. This is precisely the intent of the correlated prior method: the smooth features in w(z) are constrained by the data, with no bias induced by the prior, while the high-frequency features are constrained by the prior. Thus, our w(z)CDM model effectively has only four additional degrees of freedom compared with Λ CDM.

Dependence of the result on the correlated prior. To investigate the dependence of our result on the strength of the correlated prior, in Fig. 3, we plot the Bayes factor and the statistical significance of $w \neq -1$ as a function of σ_D , which is a parameter that is added to the inverse covariance matrix of the correlated prior effectively to strengthen it. We find that neither the ALL12 nor the ALL16 dataset provides evidence for a dynamical dark energy at all prior strengths, although the Bayes factor for ALL16 is generally much less negative than that of ALL12 for all priors. For example, the Bayes factor grows from -6.7 ± 0.3 to -3.3 ± 0.3 using ALL12 and ALL16 respectively, which corresponds to $\sigma_D = 3$. Conversely, the plot shows that, if the w(z)CDM model was true, DESI++ would be able to provide a decisive Bayesian evidence over a wide range of prior strengths.

Data availability. The data that support the plots within this paper and other findings of this study are available from the corresponding author upon reasonable request.

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Author contributions

G.-B.Z. proposed the idea, performed the dark energy reconstruction, evidence calculation, principal-component analysis and tension calculation. M.R. and Y.W. contributed to the tension calculation. G.-B.Z. and L.P. wrote the draft, and all other co-authors commented on and helped to improve the manuscript and/or contributed to the BOSS data analysis.

Competing interests

The authors declare no competing financial interests.

Additional information

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